

**Adjusting Self-Assessed Health for Potential Bias**  
**Using a Random-Effects Generalized Ordered Probit model**

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The authors are Ph.D. Graduate Research Assistant and Professor, respectively, in the School of Economic Sciences, Washington State University. This research uses data from China Health and Nutrition Survey (CHNS). We thank the National Institute of Nutrition and Food Safety, China Center for Disease Control and Prevention, Carolina Population Center, the University of North Carolina at Chapel Hill, the NIH (R01-HD30880, DK056350, and R01-HD38700) and the Fogarty International Center, NIH for financial support for the CHNS data collection and analysis files from 1989 to 2006 and both parties plus the China-Japan Friendship Hospital, Ministry of Health for support for CHNS 2009 and future surveys.

## **Abstract**

In this paper we assess how socioeconomic conditions affect self-assessed health and objective measure of health. Using a random-effects generalized ordered probit model with data from China Health and Nutrition Survey, we test for heterogeneity by socioeconomic status. The results show that individuals with high income relative to a comparison group are less likely to self-assess poor health; they are more likely to report good health; but they are not more likely to report extremely good health. These socioeconomic variables do not significantly affect an objective measure of health. Although SAH captures many aspects of health elements, it might be biased on some socioeconomic features.

**Keywords:** SAH, Reporting Heterogeneity, relative income

JEL classification: I12, I14, C25

# 1. Introduction

Self-assessed health (SAH) is a commonly used measure of individual health in a wide range of policy studies. It is often used to analyze how health responds to lifestyle and policy, as well as in distributional studies (Contoyannis and Jones 2004, Balia and Jones 2008, Costa-Font, et al. 2013). But it is often asked how well SAH adequately measures true health.

There is some evidence that SAH may be malleable depending on the survey method. Crossley and Kennedy (2002), using data from the Australian National Health Survey show that 28% of respondents in a random sub-sample which was surveyed twice changed their SAH level after giving answers to additional health related questions. Clarke and Ryan (2006) found a similar variation when SAH was again asked twice of respondents (the first in a personal interview and second in a self-completion survey). Greene et al. (2014) note an inflation of SAH. They found that “the overwhelming majority of responses fall in either the middle category or the one immediately to (its) ‘right’” and such responses are more favorable than should be expected given more objective medical indicators.

Because SAH is a subjective reporting index, there is also an immediate concern about heterogeneity in reporting. Shmueli (2003) shows extensive reporting heterogeneity in SAH that depends on a large number of socioeconomic factors, including income. Vaillant and Wolfe (2012) find the difference between SAH and objective measures is more pronounced between individuals than it is within individuals over time. One possible explanation for socioeconomic related heterogeneity is a difference in reference groups or points, depending on their demographic and social-economic characteristics (Kerkhofs and Lindeboom, 1995; Lindeboom and Van Doorslaer, 2004). Lindeboom and Van Doorslaer (2004) proposes a test for differential

reporting in ordered response models which enables to distinguish between cut-point shift and index shift using Canadian National Population Health Survey data. They find clear evidence of index shifting and cut-point shifting for age and gender, but not for income, education or language.

The hypothesis underlying the present paper is that individuals' assessment of their own health may depend on one's relative condition in one's subgroup. In research about happiness, Easterlin (1974, 1995) argues that within a country at a given time those with higher incomes are, on average, happier. However, raising the incomes of all does not increase the happiness of all because it is relative income not absolute income which affects happiness. We believe that SAH may have a similar relationship, where the comparison group for an individual might be defined by a localized reference group. To the extent socioeconomic variables like ethnicity and income determine a localized reference group, they would therefore affect SAH, an idea propagated in Wilkinson (1997). Raising the income of all may increase the health of all, but may not increase the SAH of all, especially for those with high income level.

In the research cited above, most papers use Ordered Probit or Logit models, assuming that the coefficients of independent variables do not vary between categories of the dependent variable. This assumption conceals possible heterogeneous effects of some independent variables. In addition, none use relative socioeconomic status in the regression. To fill these gaps in literature, we use a Random-Effects Generalized Ordered Probit Model (Pfarr et al., 2011), to identify the correlation with SAH and how the cut-points in assessing health vary with socioeconomic factors. Most specifically, we are interested in how relative income influences self-assessed health status. We also test for bias in SAH by using an objective health measure as the dependent variable in an otherwise identical model. Most socioeconomic variables affect people's SAH, but not affect

the objective health measure. Hence, this provides evidence that these variables cause a bias in SAH.

The rest of the paper is organized as follows. Section 2 introduces the framework of the random effect generalized ordered probit model; Section 3 introduces the data and variables we use in the analysis. The results are discussed in part 4, and part 5 offers conclusions.

## 2. The Empirical Framework

The World Health Organization (WHO) defines health as “a state of complete physical, mental and social well-being and not merely the absence of disease or infirmity” (WHO, [www.who.int/about/definition/en/print.html](http://www.who.int/about/definition/en/print.html)). Objective measures of health usually focus on disease or infirmity (the part of this definition that WHO categorically rejects as a whole measure of health) building functional indices founded on diagnostic, prognostic, and evaluative criterion (McDowell, 2006) or the incidence or absence of specific ailments. SAH, on the other hand, is more abstractly defined, with individuals asked to assign themselves to discrete categories that range from poor to excellent, often without much guidance. Underlying both objective measures of health and SAH is true health. Because one component of true health is the presence or absence of disease, it is likely that when people assess their health some objective measures of health go into that assessment. The random effects generalized ordered probit model that follows takes such behavior into account (Pfarr et al., 2010, 2011).

True health  $H_{it}^*$ , individual  $i$ 's health status in time  $t$ , is an unobserved latent variable governed by the equation

$$H_{it}^* = \alpha_i + X_{it}' \beta + \varepsilon_{it}, \varepsilon_{it} \sim N(0,1)$$

Where  $X_{it}'$  is a vector of independent variables which help determine true health. In the random effects panel data model  $\alpha_i$  represents an individual effect with a zero mean and variance  $\sigma^2$ , so  $\rho = \sigma^2 / (1 + \sigma^2)$  is the share of total variability in  $H_{it}^*$ , attributable to the individual effect. The vector  $\beta$  are parameters and  $\varepsilon_{it}$  is a random term independent of individual characteristics. Included in the vector of independent variables are individuals' demographic and socio-economic features, lifestyle, genetic disposition, current ailments and diseases, and luck. Let  $H_{it}^S$  be self-assessed health (SAH), usually obtained by survey. People are asked a question like "What do you think about your health status", choosing from a numerical scale to represent poor, fair, good and excellent health. In our data, SAH is given by a four point scale. We assume underlying the regression is the following decision;

$$\begin{aligned} H_{it}^S = 1 &\leftrightarrow H_{it}^* \leq \mu_{i1} \\ H_{it}^S = j &\leftrightarrow \mu_{ij-1} < H_{it}^* \leq \mu_{ij}, j = 2, 3 \\ H_{it}^S = 4 &\leftrightarrow H_{it}^* > \mu_{i3} \\ \mu_{ij} &= \mu_j + z_i' \gamma_j \end{aligned} \quad (1)$$

$$(2)$$

which is a form of censoring. The  $\mu_{ij}$ 's are unknown individual specific parameters to be estimated with  $\beta$ .

With four categories we have three thresholds;  $\mu_{i1} = 0$ ,  $\mu_{i2} = \mu_2 + z_i' \gamma_2$ ,  $\mu_{i3} = \mu_3 + z_i' \gamma_3$  where  $\gamma_2$  and  $\gamma_3$  are parameters to be estimated and  $z_i$  is a subset of  $X_{it}$ . The model is equivalent to three binary logistic regressions where categories of the dependent variables are combined; to find  $\mu_{i1}$  category  $H_{it}^S = 1$  is contrasted against categories  $H_{it}^S = 2,3,4$ ; for  $\mu_{i2}$  categories

$H_{it}^S = 1, 2$  are contrasted with  $H_{it}^S = 3, 4$ ; and to find  $\mu_{i3}$  categories  $H_{it}^S = 1, 2, 3$  are contrasted against category  $H_{it}^S = 4$  (Williams 2006). If  $\gamma_2$  and  $\gamma_3$  are nonzero, the thresholds are conditional on  $z_i$ , unlike the normal probit model where the thresholds are the same for all individuals.<sup>1</sup> Hence a generalized ordered probit model accounts for individual heterogeneity through the thresholds.<sup>2</sup> Imposing our functional forms for the thresholds we have

$$\begin{aligned}
H_{it}^S &= 1 \text{ if } H_{it}^* \leq 0 \\
H_{it}^S &= 2 \text{ if } 0 \leq H_{it}^* \leq \mu_2 + z_i' \gamma_2 \\
H_{it}^S &= 3 \text{ if } \mu_2 + z_i' \gamma_2 \leq H_{it}^* \leq \mu_3 + z_i' \gamma_3 \\
H_{it}^S &= 4 \text{ if } H_{it}^* \geq \mu_3 + z_i' \gamma_3
\end{aligned}$$

which gives the following probabilities

$$\begin{aligned}
P_1 &= \text{Prob}(H_{it}^S = 1 | X_i, Z_{it}) = F(-\alpha_i - X_{it}' \beta) \\
P_2 &= \text{Prob}(H_{it}^S = 2 | X_{it}', Z_{it}) = F(\mu_2 + z_{it}' \gamma_2 - (\alpha_i + X_{it}' \beta)) - F(-\alpha_i - X_{it}' \beta) \\
P_3 &= \text{Prob}(H_{it}^S = 3 | X_{it}', Z_{it}) = F(\mu_3 + z_{it}' \gamma_3 - (\alpha_i + X_{it}' \beta)) - F(\mu_2 + z_{it}' \gamma_2 - (\alpha_i + X_{it}' \beta)) \\
P_4 &= \text{Prob}(H_{it}^S = 4 | X_{it}', Z_{it}) = 1 - F(\mu_3 + z_{it}' \gamma_3 - (\alpha_i + X_{it}' \beta))
\end{aligned}$$

We use MLE and a corresponding log-likelihood function

$$\begin{aligned}
\ln L &= \sum_{H_{it}^S=1} F(-\alpha_i - X_{it}' \beta) + \sum_{H_{it}^S=2} [F(\mu_2 + Z_{it}' \gamma_2 - (\alpha_i + X_{it}' \beta)) - F(-\alpha_i - X_{it}' \beta)] \\
&+ \sum_{H_{it}^S=3} [F(\mu_3 + Z_{it}' \gamma_3 - (\alpha_i + X_{it}' \beta)) - F(\mu_2 + Z_{it}' \gamma_2 - (\alpha_i + X_{it}' \beta))] \\
&+ \sum_{H_{it}^S=4} [1 - F(\mu_3 + Z_{it}' \gamma_3 - (\alpha_i + X_{it}' \beta))]
\end{aligned}$$

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<sup>1</sup> The traditional ordered probit assumes the categories are “parallel” and differ only by the intercept. The generalized ordered probit does not impose this assumption, which is often violated in practice.

<sup>2</sup> It is common to report the results from Generalized Ordered Probit as (in our case) three different sets of estimates that include the thresholds in the estimates of  $\beta$  and then separately report the values of the  $\gamma$ . This is how we report our results in Tables 4A and 4B below.

### **3. Data**

We use the data from China Health and Nutrition Survey (CHNS), which is an international collaborative project between the Carolina Population Center at the University of North Carolina at Chapel Hill and the National Institute of Nutrition and Food Safety at the Chinese Center for Disease Control and Prevention. This survey was conducted in nine provinces in China for nine waves from year 1989 to year 2011. Among the dataset, there are 4 years with reported individual's self-assessed health (1997, 2000, 2004 and 2006). Since some individuals were not surveyed every year, we use only those observations that have at least 3 years of data. After data cleaning, the effective unbalanced panel includes 22055 observations. Among them, 4665 observations are in year 1997, 4983 observations are in year 2000, 6401 observations are in year 2004, and 5997 observations are in year 2006.

#### **3.1 Variables: A production function for health**

We follow the theoretic framework in Contoyannis and Jones (2004) to choose variables for equations (1) and (2). Table 1 below shows the variables we include. For analytical purposes, we divided the variables into groups representing health behaviors, objective health measures, education, marital status, work status, physical and regional variables. Relative health was kept as its own group.

Health behaviors include variables that measure sleep, smoking, habits on alcohol consumption and exercise. Sleep is a dummy variable which takes a value 1 if an individual sleep 7 to 9 hours and takes value 0 otherwise. For smoking, we divide people into three kinds, current smoker, previous smoker and people who never smoked. Current smoker is the excluded category. We use two variables to indicate alcohol consumption, "Alcohol\_freq" and "Alcohol\_occa". People



who don't drink alcohol at all are excluded. The "Exercise" variable takes value 1 if the person participates at least one kind of exercise. The exercises in the survey included Kung Fu, Gymnastics, dancing, acrobatics, Track and field (running, etc.), swimming, Soccer, basketball, tennis, Badminton, volleyball and others. "Nobese" means the individual is *not* obese.

For the objective health measures, the survey asked respondents if a doctor had ever told them they had one of five conditions, high blood pressure, diabetes, myocardial infarction, Apoplexy, and Fracture. These chronic diseases were chosen because they affect people's life quality in many ways. Although they do not cause death in a short time, they are leading causes of death in the long run.

*(Insert Table 1. Independent variable)*

Relative income is often considered a substitute for social class (Contoyannis and Jones 2004; Wilkinson 1997). We hypothesize that people self-refer on specific socioeconomic features, for example, comparing themselves to others of the same education level. Here we use relative income in the same province<sup>3</sup> and education level as two sources of self-reference. The relative income of individual  $i$ , who lives in province  $j$  (or with education level  $j$ ), at year  $t$  is calculated by the equation

$$rltv\_income_{ijt} = \frac{hhinc_{it}}{E(hhinc)_{jt}}$$

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<sup>3</sup> We tried to use relative income within a respondent's town, but that provided insufficient variability as incomes do not vary much within towns. Moreover, we test, therefore, if people compare not just within their own community, but also to nearby communities. "Province" in China is like the concept of "State" in the US.

where  $hhinc_{it}$  is the household income of individual  $i$  at time  $t$ , and  $E(hhinc)_{jt}$  is the average household income in province  $j$  (or with education level  $j$ ) at year  $t$ . When relative income equals to 1 it means that income is at the average level. Individuals with income above average have a relative income great than 1; those with an income below the average have a relative income less than 1.

Most other grouped variables are self-explanatory except for “Urban\_hukou”. Hukou is a special concept in China for household registration. China has two kinds of Hukou that distinguish people who live in city or urban area from people who live in rural area. Urban\_hukou=1 indicates the respondent is registered in an urban area.

### **3.2 Descriptive analysis**

Table 2A presents the mean values of the variables by the four SAH subgroups. The subgroup reporting SAH=1 feel their health status is “poor”. SAH=2 means health level is “fair”; SAH=3 means health level is “good”; SAH=4 indicates health level “excellent”.

Relative income, both refer to the same province and same education level is highly related to SAH status. Both “good” health and “excellent” health subgroups have above average incomes. People who assess their health as poor have income significantly lower than the average. However, the difference between the excellent and good health subgroups is less significant than the difference between other SAH subgroups.

Among the behavior variables, sleep has an ambiguous trend among the four SAH subgroups, while exercise has a clear increasing trend from unhealthy to healthy subgroups. From the exercise and habits variables, we see people who feel healthy have a better habits and exercise

more. The poor-health subgroup has a higher proportion of non-smokers and former smokers. People in the healthier subgroups have a lower rate of obesity.

Objective healthy measures are highly consistent with people's SAH. People in the healthy SAH subgroups have lower morbidity rates of all the diseases we use. Especially for the excellent health subgroup, few people are diagnosed of those severe and chronic diseases. Individual's average number of illness decreases from poor health group to excellent health group.

*(Insert Table 2A. Means of the variables by SAH subgroups)*

People with more education tend to report higher levels of health. For example, the proportion of individuals with middle school, high school and college or university degree (or higher) increases as we move from unhealthy to healthy. However, a higher proportion of divorce and separation are observed in fair and good subgroups. The proportion of single people is larger in higher health levels.

A higher proportion of unemployed, house keeper, disable and retired people are observed in the "poor" health subgroup. In addition, the rate of unemployed in subgroup SAH=1 is much higher than that in other subgroups. The proportion of people doing agricultural labor work is higher in "poor" and "fair" health subgroup.

Physical condition and living conditions also have a clear pattern. Those indicating they have excellent health are more likely to be male, younger and taller. And those indicating poor health and excellent health status are more likely to live in the urban areas.

*(Insert Table 2B. Means of the variables by Objective Health)*

We then define an objective measure of health to compare with the SAH. As an objective measure of health we used the number of severe or chronic diseases (denote as “Num\_diseases”) for which an individual has been diagnosed. The list of possible diseases includes high blood pressure, diabetes, myocardial infarction, apoplexy, and fracture. To be consistent with the SAH scale, “Num\_diseases=1” represents for the most unhealthy group, which means the person has more than one diseases; “Num\_diseases=2” represents that the person has one of these diseases; “Num\_diseases=3” is the healthiest group, which means the person is diagnosed with none of these diseases.

Table 2B shows the mean of variables by “Num\_diseases”. Mean of the relative income doesn’t change apparently. The healthiest group (Num\_diseases=3, No disease) even has the relative income slightly lower than the other two groups. A similar trend can be see regard to the education variable. The proportion of people with lower education level (elementary school and middle school degrees) increases when we move from the unhealthy group to healthy group, while the proportion of observations with higher education level (high school and college or above) decreases.

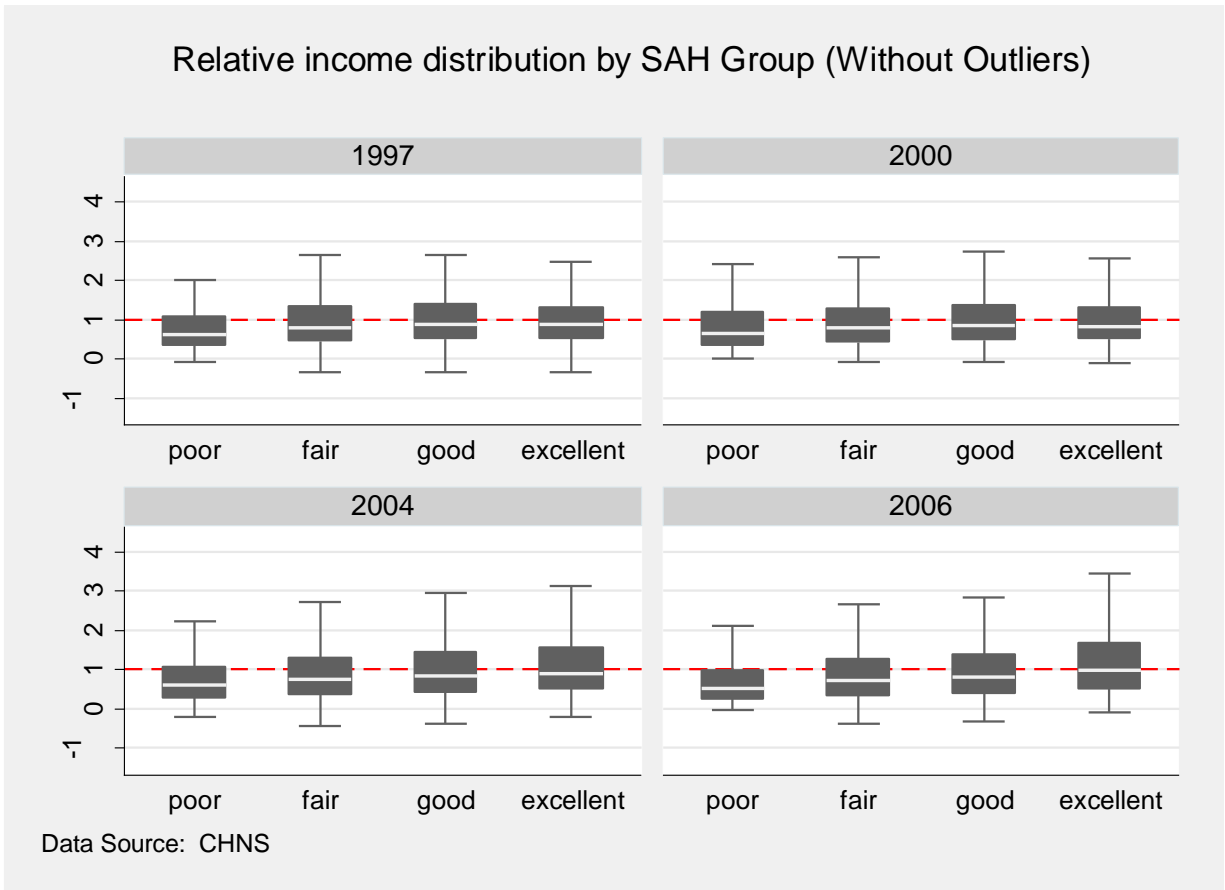


Figure 1. Relative household income distribution by SAH  
(Without outliers greater than 1.5 times the 75<sup>th</sup> percentile)

Figure 1 shows the relative household income distribution by SAH subgroups in different years. The line in the box is the medium of the relative income of every subgroup, while the boxes represent the portion between the 25th percentile and the 75th percentile. In 1997, the box for the poor health subgroup is below the dashed line. It means most people who report poor health earn incomes below average. We expected that wealthier people would be healthier, but the excellent health subgroup has a lower relative income when compared to the good health subgroup. The same situation is observed in year 2000, but in 2004 and 2006 the healthiest subgroup has the largest proportion of observations with a high relative household income. Compared to the other

categories of SAH, the boxes for the poor health subgroup are comparatively narrow, indicating the variation of relative income in this subgroup is less than in the other subgroups.

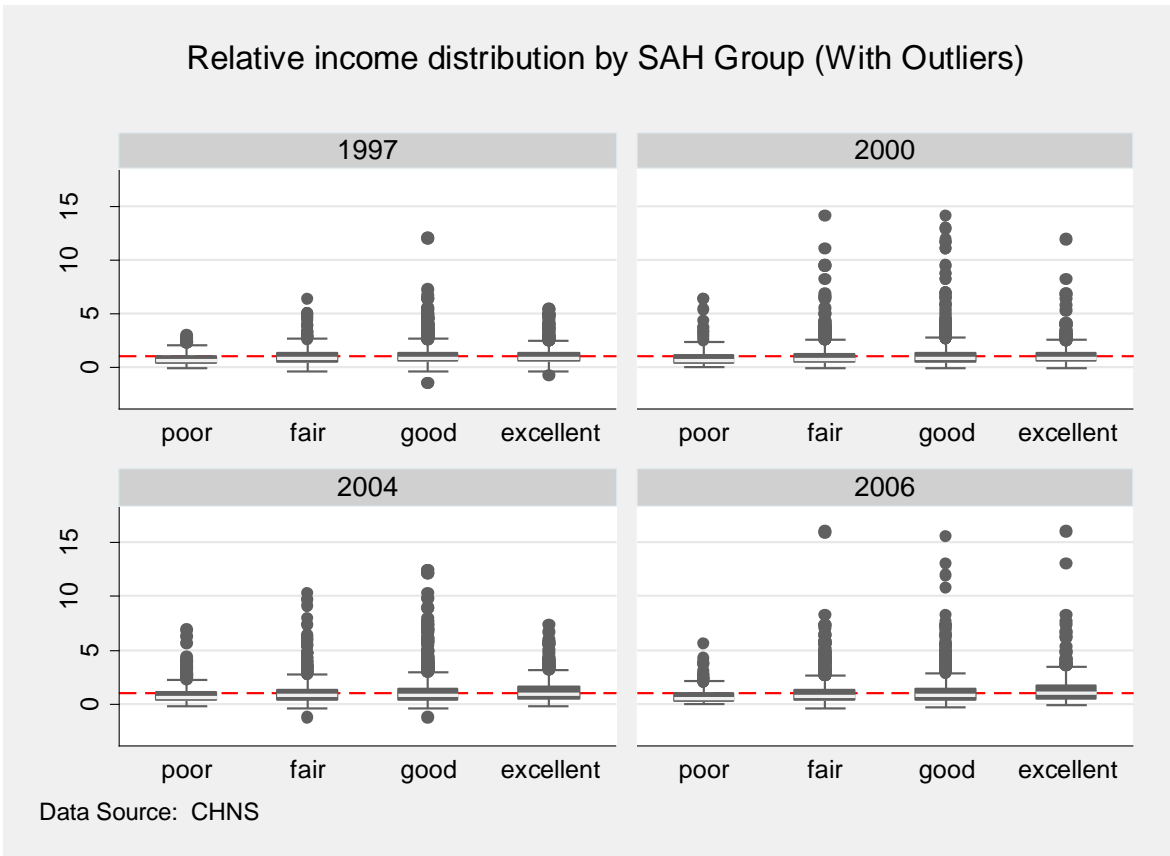


Figure 1. Relative household income distribution by SAH (with outliers greater than 1.5 times the 75<sup>th</sup> percentile)

Figure 2 includes Outsider comparing the distribution of relative income by SAH group. The fair and good health subgroups show a large spread of relative income, with high relative income cluster in these two subgroups. The good health subgroup shows the largest spread.

Table 3 shows the correlation between the number of diagnosed illness and self-assessed health disaggregated by relative income. People diagnosed with more kinds of diseases are less likely to report good health, although the magnitude of the correlation is small. We note that the correlation becomes strong when relative income increases.

*(Insert Table 3. Correlation between number of illness and SAH)*

## **4. Results**

### **4.1 SAH as measure of health**

Table 4A shows the results from two different regression models, a random effects ordered probit and a random effects generalized ordered probit. SAH is used as a measure of health. As we can see from table 4A, relative income, sleep, education degree of middle and high school, single, widow, unemployment, disable, Urban\_Hukou, male, height, diagnosed of hypertension and apoplexy have different coefficients in the three parts of the generalized ordered probit model, i.e., these variables violate the parallel line assumption. Table 4B show the coefficient  $\gamma_2$  and  $\gamma_3$  of these variables derived from the estimates in Table 4A.

*(Insert Tables 4A and 4B. Random effect ordered probit and generalized ordered probit models)*

We first pay attention to the variables that satisfy the parallel line assumption. The two smoking behavior variables have opposite effect; people who quit smoking are more likely to report poor health, while people who never smoke do not have a significant difference from current smokers.<sup>4</sup> The two alcohol behavior variables also have opposite sign coefficients, although only frequent use is significant at conventional levels. Frequent alcohol users report good health status. People usually doing exercise report better health than people who do not. Generally speaking, people with more education report they are healthier. Divorced and separated people tend to report poorer health than the base group, people who are married. Among the variables about

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<sup>4</sup> Poor health may lead people to quit smoking, creating an potential endogeneity problem with this variable that needs further exploration.

working status, people unemployed, involved in housekeeping and disabled, people who shift work are all have worse self-assessed health than people work normally.

Of primary interest are those variables that violate the parallel assumption, especially relative income. The result suggests that those who have higher relative income tend to report better health. The positive effect of a relative income is especially high among those who report themselves to be poor health as opposed to fair, good, or excellent health. When translated to the  $\gamma$  coefficient (table 4B) it indicates that relative income lowers the threshold that pushes an individual to the next highest level of SAH, so those with higher relative income are more likely to be in the next highest category of SAH, indicating that those with higher relative income are more likely to say their health is better if their SAH is in the fair or good categories. Relative income is not statistically significant in the excellent health category. In sum, these results indicate that having high relative income lowers the probability that people will self-assess their health as poor, but also does not increase the probability that they will assess their health as excellent.

Educational attainment may be another way people self-reference. Those who have a high-school degree are unlikely to say they are in poor health, but are also unlikely to say they are in excellent health. The estimates on good sleeping behavior are interesting. Generally, sleeping well has a positive effect for the individual to choose fair or good health against poor health level, but they are also unlikely to choose excellent health level compare to good, fair and poor. Living in an urban area increases the probability of feeling extremely healthy. Compared to females, males are more conservative about assessing good health -- they concentrate in the two middle categories. Tall people self-assess healthier, and the effect becomes stronger when health level increases. All the disease variables make people feel unhealthy. We also report  $\rho$  for both



models in Table 4A. In both models, about 22 percent of total variation in SAH can be attributed to individual fixed effects. This translates to a variance of about 0.282 for  $\alpha_i$ .

*(Insert Table 5. Random effect generalized ordered probit model with variable relative BMI)*

As indicated, SAH is affected by the person's objective health. Table 5 shows the results adding variable relative BMI to the model. Even though relative BMI is significant with a p-value < 0.001, the signs and significance of the other variables is consistent with the earlier results.

*(Insert Table 6. Marginal effect of random effect generalized ordered probit model)*

Table 6 provides the marginal effect of the random effect generalized ordered probit model. When relative income increases by 1 unit, the probability of reporting poor health decreases by about 1% while the probability of reporting good health increases by 1.19%. In our data, the highest relative income is about 16 (a value of 1 means the respondent earns an average income). At that level the probability of reporting poor health is decreased by 15%, and the probability of reporting good health is increased by 15%. Education, another important socioeconomic variable, also increases the probability of people reporting good health. Attaining a high school, technical or vocational degree increases the probability of reporting good health by approximately 5%.

## **4.2 Objective diseases as a measure of health**

In this section, we discuss the model using a more objective measure of health as a dependent variable.

*(Insert Table 7. Random effect generalized ordered probit model with objective health measure)*

Table 7 shows the regression results when we use the number of disease as the measure of health. The variable of most interest is relative income. As we can see from table 7, having a relatively high income does not affect the possibility of having severe diseases. More education, which positively affected people self-assessed health, healthier, does not reduce the incidence of disease. In fact, we see an opposite impact -- people who graduate from elementary school and high school are more likely to be diagnosed with a disease compared to the illiterate group. Obesity is also a risk factor. As we found with SAH, good sleep habits improves health. Finally, working in agriculture lowers the risk of having one or more of these diseases, lowers objective risk, even though those working in agriculture have lower SAH. Table 8 shows the marginal effect of random effect Generalized Ordered probit model with objective health measure. When relative income increases by 1, the probability of getting one or more severe diseases is not affected.

*(Insert Table 8. Marginal effect of random effect Generalized Ordered probit model with objective health measure)*

### **4.3 Relative income by education levels**

Besides considering the relative income referring to the people in the same province, we also calculate a relative income refer to the same education level subgroup. The relative income referring to the same education level is used for a robust check for the model. It is used for several reasons. First, people usually compete with others in the same education level. Second, when comparing the income level with others, we are usually more sensitive to the group of the same education level. Third, people with same education level may have a similar consumption habit, which leads to a similar expenditure pattern.

Table 9 and 10 show the regressions when we use the relative income refer to the same education level subgroups. We can see the results are very consistent with regression in table 5 and 7 except tiny coefficient value difference. Being relatively richer within the same education level group makes it easier for the individual to feel “fair” and “good” health status, but it doesn’t help for people to feel extremely healthy. When using a more objective measure of health, this relative income has no significant effect to make people healthier.

## **5. Conclusion**

We use a random effect generalized ordered probit model to test for individual heterogeneity in self-assessed health. While several variables contribute to such heterogeneity, we focus on the influence of relative household income. And we also compare it with the model using number of diseases as an objective health measure. Using data from the China Health and Nutrition Survey (CHNS), In general, SAH is not only affected by individual’s physical state, like age, height, and diseases, but also significantly affected by socioeconomic status, like education level, working situation, marriage status and income level refer to comparable groups. We find that people with high relative income feel better about their health and, more importantly, they have a lower threshold to assess that they have good health. People with high relative income are less likely to report poor health, but they are also less likely to report extremely healthy. However, most socioeconomic variables, like relative income and education level, doesn’t influence the objective measure of health, which is generally accepted to be closer to people’s “true health”.

The results imply that no matter how the individual compares their income with others, regionally or by education level, relative higher income gives them an optimistic feeling

regarding their health status, though they may not actually healthier than those with lower relative income. We should be careful when using SAH as a measurement of health in research, especially when we study the relationship between economic inequality and health. Although SAH capture many aspects of health elements, it might be biased on some socioeconomic features. The results of this study might raise more discussion about bias in SAH and how to adjust SAH as a measurement of individual health in economic and policy research.

## References

- Balia, S. and A. M. Jones (2008). "Mortality, lifestyle and socio-economic status." *Journal of Health Economics* 27(1): 1-26.
- Clarke, P. M. and C. Ryan (2006). "Self-reported health: reliability and consequences for health inequality measurement." *Health Economics* 15(6): 645-652.
- Contoyannis, P. and A. M. Jones (2004). "Socio-economic status, health and lifestyle." *Journal of Health Economics* 23(5): 965-995.
- Costa-Font, Joan and Hernandez Quevedo, Cristina and Sato, Azusa, A 'Health Kuznets' Curve'? Cross-Country and Longitudinal Evidence (October 31, 2013). *CESifo Working Paper Series* No. 4446. Available at SSRN: <http://ssrn.com/abstract=2348070>
- Crossley, T. F. and S. Kennedy (2002). "The reliability of self-assessed health status." *Journal of Health Economics* 21(4): 643-658.
- Easterlin, R. A. (1974). "Does economic growth improve the human lot? Some empirical evidence." *Nations and households in economic growth* 89: 89-125.
- Easterlin, R. A. (1995). "Will raising the incomes of all increase the happiness of all?" *Journal of Economic Behavior & Organization* 27(1): 35-47.
- Greene, William H. and Harris, Mark N. and Hollingsworth, Bruce, Inflated Responses in Measures of Self-Assessed Health (May 2014). *NYU Working Paper* No. 2451/33696. Available at SSRN: <http://ssrn.com/abstract=2443781>
- Kerkhofs, M. and M. Lindeboom (1995). "Subjective health measures and state dependent reporting errors." *Health Economics* 4(3): 221-235.
- Lindeboom, M. and E. van Doorslaer (2004). "Cut-point shift and index shift in self-reported health." *Journal of Health Economics* 23(6): 1083-1099.
- McDowell, Ian (2006). *Measuring Health: A Guide to Rating Scales and Questionnaires* (Third Edition). New York: Oxford University Press, Inc.
- Pfarr, C., Schmid, A. and Schneider, U. (2011). "Estimating ordered categorical variables using panel data: A generalised ordered probit model with an autofit procedure." *Journal of Economics and Econometrics* 54 (1): 7-23.

- Schneider, U., Pfarr, C., Schneider, B., Ulrich, V., (2012). "I feel good! Gender differences and reporting heterogeneity in self-assessed health," *The European Journal of Health Economics*, Springer, vol. 13(3), pages 251-265, June.
- Shmueli, A. (2003). "Socio-economic and demographic variation in health and in its measures: the issue of reporting heterogeneity." *Social Science & Medicine* 57(1): 125-134.
- Vaillant, N. and F.-C. Wolff (2012). "On the reliability of self-reported health: Evidence from Albanian data." *Journal of Epidemiology and Global Health* 2(2): 83-98.
- Wilkinson, R. G. (1997). "Socioeconomic determinants of health. Health inequalities: relative or absolute material standards?" *British Medical Journal* 314(7080): 591.

Table 1. Independent variable

<b>Variable</b>	<b>discription</b>
SAH	Self-assessed health
rltv_income	Household net income relative to the average income in the province
rltv_income2	Household net income relative to the average income of same education level group
<b>behavior</b>	
Sleep	1 if sleep time is between 7 and 9 hours a day, otherwise set 0
Nonsmoker	1 if the person never smoke
Smokerquit	1 if the person smoked before but quit now
Alcohol_freq	1 if have alcohol more than once or twice a week
Alcohol_occa	1 if have alcohol less than once or twice a month
Exercise	1 if the person participate at least one kind of outdoor exercise
<b>Objective</b>	
Hyper	1 if the person is diagnosed of high blood tension
Diabetes	1 if the person is diagnosed of diabetes
MI	1 if the person is diagnosed of myocardial infarction
Apoplexy	1 if the person is diagnosed of apoplexy
Fracture	1 if the person has a history of bone fracture
<b>Work</b>	
unemp	1 if the person is totally unemployed
housekeep	1 if the person is unemployed but is a housekeep
disable	1 if the person is unemployed because he is disable
retired	1 if the person is retired
rehired	1 if the person is rehired after retired
Work shift	1 if the person change works after 2004
Ag_labor	1 if the person participate in one or more agricultural labor work
<b>Education</b>	
Educ_1	Highest level is elementary school
Educ_2	Highest level attained is middle school degree
Educ_3	Highest level attained is high school or technical or vocational degree
Educ_4a	Highest level attained is college and university or above
<b>Marital status</b>	
Single	1 if single and never married
Divorced	1 if get divorced
Widow	1 if the spouse died
Separate	1 if Separate
<b>physical</b>	
Male	1 if the person is male

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Height

Age

**Region**

Urban\_hukou

1 if the person's "hukou" is urban

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Table 2A. Means of the variables by SAH Subgroups

Variable	SAH=1 (obs=1341)	SAH=2 (obs=6905)	SAH=3 (obs=10999)	SAH=4 (obs=2810)
rltv_income1 (Province)	0.8009929	0.9944218	1.092314	1.133842
rltv_income2 (Education)	0.7819337	0.9426062	0.9984135	1.030518
<b>Behavior</b>				
sleep	0.4198359	0.4764663	0.4164924	0.4320285
exercise	0.0574198	0.0855902	0.0981907	0.1241993
nonsmoker	0.7136465	0.702824	0.6626057	0.6327402
smokerquit	0.049217	0.0291093	0.0199109	0.0185053
alcoholfreq	0.1469053	0.2152064	0.2701155	0.3160142
alcohol_occa	0.0618941	0.0844316	0.0980998	0.0903915
nobese	0.9261745	0.939609	0.951541	0.9409253
<b>Objective</b>				
hyper	0.2281879	0.1229544	0.0466406	0.0270463
diabete	0.0611484	0.0196959	0.0069097	0.0017794
MI	0.0208955	0.0088444	0.0010926	0.0003561
apoplexy	0.0656227	0.0081101	0.0022729	0.0003559
fracture	0.0805369	0.0544533	0.0307301	0.016726
ill_num	0.4563758	0.2140478	0.0876443	0.0462633
<b>Work</b>				
unemp	0.1327368	0.0734251	0.0598236	0.0715302
housekeep	0.2013423	0.1562636	0.1017365	0.0814947
disable	0.0298285	0.0034757	0.0012728	0.0014235
retired	0.1700224	0.1338161	0.0776434	0.058363
rehired	0.00522	0.0060825	0.0084553	0.0081851
workshift	0.0290828	0.039971	0.0473679	0.0483986
ag_labor	0.284862	0.2912382	0.2309301	0.2053381
<b>Education</b>				
educ_1	0.2334079	0.2377987	0.2304755	0.213879
educ_2	0.1715138	0.24895	0.3173925	0.3548043
educ_3	0.1096197	0.1452571	0.1942904	0.2252669
educ_4a	0.01566	0.0267922	0.0337303	0.0466192
<b>Marital status</b>				
single	0.0350485	0.0457639	0.0695518	0.1160142
divorce	0.0067114	0.0098479	0.0084553	0.005694
widow	0.1342282	0.087328	0.0499136	0.0209964



separate	0.0014914	0.0017379	0.0018183	0.0007117
<b>Physical</b>				
age	56.52573	51.64374	45.13801	40.96192
BMI	22.60845	22.8279	22.97065	23.37462
height	157.3863	158.7348	160.6702	163.1306
male	0.3907532	0.422882	0.4969543	0.552669
<b>Region</b>				
urban_hukou	0.2923192	0.2773353	0.2614783	0.3014235

Table 2B. Means of the variables by Objective Health

variable	Num_diseases=1 (Multi-diseases)	Num_diseases=2 (One disease)	Num_diseases=3 (No disease)
rltv_income	1.086426	1.084459	1.0442
rltv_income2	1.091354	1.051283	0.9597798
<b>Behavior</b>			
sleep	0.4201031	0.4775042	0.4329139
exercise	0.1262887	0.1222411	0.0911398
nonsmoker	0.693299	0.6689304	0.6747967
smokerquit	0.0850515	0.0522071	0.0197815
alcoholfreq	0.2216495	0.2567912	0.251204
alcohol_occa	0.0670103	0.0772496	0.0927451
nobese	0.8402062	0.8633277	0.9569675
<b>Work</b>			
unemp	0.0902062	0.0785229	0.068562
housekeep	0.1340206	0.1612903	0.1172907
disable	0.0231959	0.0101868	0.0025374
retired	0.4742268	0.2580645	0.0713583
rehired	0.0257732	0.012309	0.0065248
workshift	0.0180412	0.0339559	0.0458288
ag_labor	0.128866	0.2245331	0.2553467
<b>Education</b>			
educ_1	0.193299	0.229202	0.231785
educ_2	0.1881443	0.2270798	0.3018487
educ_3	0.1881443	0.1672326	0.17881
educ_4a	0.0386598	0.0420204	0.0307597

<b>Marital status</b>				
single	0.0206186		0.033107	0.0708405
divorce	0.007732		0.0067912	0.0086479
widow	0.1469072		0.1260611	0.0537
separate	0		0.0008489	0.0017607
<b>Physical</b>				
age	62.60567		56.42275	45.91963
BMI	25.22453		24.39153	22.7346
height	160.059		159.9562	160.2076
male	0.4819588		0.4787776	0.4737196
<b>Region</b>				
urban_hukou	0.5025773		0.3943124	0.2540521

Table 3. Correlation between number of illness and SAH

rltv_income	<=0.5	>0.5 & <=1	>1 & <=2	>2 & <=3	>3
corr	-0.2226	-0.2284	-0.2316	-0.2644	-0.2894

Table 4A. Random Effect Ordered Probit and Generalized Ordered Probit

	Ordered probit	Generalized Ordered probit		
		1 vs. 2-4	1-2 vs. 3-4	1-3 vs. 4
sleep	0.091*** (0.02)	0.212*** (0.04)	0.075** (0.03)	0.016 (0.04)
nonsmoker	0.007 (0.03)	0.009 (0.03)	0.009 (0.03)	0.009 (0.03)
smokerquit	-0.077 (0.06)	-0.072 (0.06)	-0.072 (0.06)	-0.072 (0.06)
alcoholfreq	0.176*** (0.02)	0.175*** (0.03)	0.175*** (0.03)	0.175*** (0.03)
alcohol_occa	-0.035 (0.03)	-0.036 (0.03)	-0.036 (0.03)	-0.036 (0.03)
exercise	0.055 (0.03)	0.056 (0.03)	0.056 (0.03)	0.056 (0.03)
nobese	-0.150*** (0.04)	-0.147*** (0.04)	-0.147*** (0.04)	-0.147*** (0.04)
educ_1	0.057* (0.03)	0.054 (0.03)	0.054 (0.03)	0.054 (0.03)
educ_2	0.108*** (0.03)	0.217*** (0.05)	0.117*** (0.03)	0.040 (0.04)
educ_3	0.141*** (0.04)	0.210*** (0.06)	0.175*** (0.04)	0.040 (0.04)
educ_4a	0.157* (0.06)	0.132* (0.06)	0.132* (0.06)	0.132* (0.06)
single	0.024 (0.04)	0.069 (0.09)	-0.049 (0.05)	0.096 (0.05)
divorce	-0.128 (0.10)	-0.116 (0.10)	-0.116 (0.10)	-0.116 (0.10)
widow	0.052 (0.04)	0.072 (0.06)	0.105* (0.05)	-0.100 (0.08)
separate	-0.122 (0.21)	-0.156 (0.21)	-0.156 (0.21)	-0.156 (0.21)
unemp	-0.142*** (0.03)	-0.355*** (0.06)	-0.134** (0.04)	0.006 (0.05)
housekeep	-0.084** (0.03)	-0.085** (0.03)	-0.085** (0.03)	-0.085** (0.03)
disable	-1.100*** (0.15)	-1.383*** (0.16)	-0.810*** (0.18)	-0.270 (0.28)
retired	-0.014 (0.04)	-0.022 (0.04)	-0.022 (0.04)	-0.022 (0.04)
rehired	0.262** (0.10)	0.263** (0.10)	0.263** (0.10)	0.263** (0.10)
workshift	-0.161***	-0.157***	-0.157***	-0.157***

	(0.04)	(0.04)	(0.04)	(0.04)
ag_labor	-0.084***	-0.082***	-0.082***	-0.082***
	(0.02)	(0.02)	(0.02)	(0.02)
urban_hukou	0.020	-0.032	-0.009	0.088**
	(0.03)	(0.04)	(0.03)	(0.03)
male	-0.029	0.014	0.034	-0.144***
	(0.03)	(0.05)	(0.04)	(0.04)
height	0.016**	0.011**	0.013**	0.024***
	(0.00)	(0.00)	(0.00)	(0.00)
age	-0.019***	-0.015***	-0.021***	-0.017***
	(0.00)	(0.00)	(0.00)	(0.00)
hyper	-0.482***	-0.429***	-0.530***	-0.381***
	(0.03)	(0.05)	(0.04)	(0.07)
diabete	-0.713***	-0.716***	-0.716***	-0.716***
	(0.08)	(0.08)	(0.08)	(0.08)
apoplexy	-1.028***	-1.190***	-0.727***	-0.854*
	(0.10)	(0.12)	(0.14)	(0.39)
fracture	-0.404***	-0.406***	-0.406***	-0.406***
	(0.04)	(0.04)	(0.04)	(0.04)
rltv_income	0.033***	0.121***	0.038***	0.002
	(0.01)	(0.02)	(0.01)	(0.01)
time	-0.065***	-0.089***	-0.097***	0.011
	(0.01)	(0.02)	(0.01)	(0.02)
_cons	0.542*	0.995*	-0.368	-4.220***
	(0.27)	(0.46)	(0.31)	(0.38)
rho	0.219***		0.220***	
	(0.01)		(0.01)	
N	22055		22055	

Standard errors in parentheses

\*  $p < 0.05$ , \*\*  $p < 0.01$ , \*\*\*  $p < 0.001$

Table 4B. Random effect ordered probit and Generalized Ordered probit model

Random Effect Generalized Ordered probit		
	gamma2	gamma3
sleep	-0.13744	-0.19625
educ_2	-0.10027	-0.17683
educ_3	-0.03429	-0.16947
single	-0.11776	0.027264
widow	0.033152	-0.17204
unemp	0.221356	0.361084
disable	0.573803	1.113395
urban_hukou	0.023093	0.120321
male	0.019478	-0.15827

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height	0.001917	0.012777
age	-0.00664	-0.00235
hyper	-0.10078	0.048416
apoplexy	0.462588	0.335501
rltv_income	-0.08239	-0.11891
time	-0.00829	0.099602
_cons	-1.36292	-5.21511

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Table 5. Random effect generalized ordered probit model with variable relative BMI

	Generalized Ordered probit (SAH)		
	1 vs. 2-4	1-2 vs. 3-4	1-3 vs. 4
sleep	0.209*** (0.04)	0.070* (0.03)	0.011 (0.04)
nonsmoker	-0.007 (0.03)	-0.007 (0.03)	-0.007 (0.03)
smokerquit	-0.082 (0.06)	-0.082 (0.06)	-0.082 (0.06)
alcoholfreq	0.164*** (0.02)	0.164*** (0.02)	0.164*** (0.02)
alcohol_occa	-0.041 (0.03)	-0.041 (0.03)	-0.041 (0.03)
exercise	0.054 (0.03)	0.054 (0.03)	0.054 (0.03)
nobese	0.155*** (0.05)	0.155*** (0.05)	0.155*** (0.05)
educ_1	0.048 (0.03)	0.048 (0.03)	0.048 (0.03)
educ_2	0.210*** (0.05)	0.108** (0.03)	0.031 (0.04)
educ_3	0.195** (0.06)	0.162*** (0.04)	0.026 (0.04)
educ_4a	0.129* (0.06)	0.129* (0.06)	0.129* (0.06)
single	0.119 (0.09)	-0.000 (0.05)	0.147** (0.05)
divorce	-0.088 (0.10)	-0.088 (0.10)	-0.088 (0.10)
widow	0.106 (0.06)	0.140** (0.05)	-0.061 (0.08)
separate	-0.143 (0.21)	-0.143 (0.21)	-0.143 (0.21)
unemp	-0.356*** (0.06)	-0.133** (0.04)	0.007 (0.05)
housekeep	-0.098*** (0.03)	-0.098*** (0.03)	-0.098*** (0.03)
disable	-1.383*** (0.16)	-0.805*** (0.18)	-0.267 (0.28)
retired	-0.049 (0.04)	-0.049 (0.04)	-0.049 (0.04)
rehired	0.232* (0.10)	0.232* (0.10)	0.232* (0.10)
workshift	-0.151***	-0.151***	-0.151***

	(0.04)	(0.04)	(0.04)
ag_labor	-0.076**	-0.076**	-0.076**
	(0.02)	(0.02)	(0.02)
urban_hukou	-0.050	-0.027	0.070*
	(0.04)	(0.03)	(0.03)
male	0.021	0.033	-0.152***
	(0.05)	(0.04)	(0.04)
height	0.010***	0.012***	0.024***
	(0.00)	(0.00)	(0.00)
age	-0.014***	-0.021***	-0.017***
	(0.00)	(0.00)	(0.00)
hyper	-0.472***	-0.572***	-0.420***
	(0.05)	(0.04)	(0.07)
diabete	-0.735***	-0.735***	-0.735***
	(0.08)	(0.08)	(0.08)
apoplexy	-1.187***	-0.712***	-0.814*
	(0.12)	(0.14)	(0.39)
fracture	-0.414***	-0.414***	-0.414***
	(0.04)	(0.04)	(0.04)
rltv_income	0.118***	0.036***	-0.000
	(0.02)	(0.01)	(0.01)
rltv_BMI	0.977***	0.977***	0.977***
	(0.08)	(0.08)	(0.08)
time	-0.095***	-0.104***	0.004
	(0.02)	(0.01)	(0.02)
_cons	-0.146	-1.526***	-5.398***
	(0.46)	(0.32)	(0.39)
rho		0.210***	
		(0.01)	
N		22055	

Table 6. Marginal effect of random effect Generalized Ordered probit model

	Marginal effects for p(SAH=1)	Marginal effects for p(SAH=2)	Marginal effects for p(SAH=3)	Marginal effects for p(SAH=4)
sleep	-0.0166*** (0.00324)	-0.00681 (0.00901)	0.0215* (0.00975)	0.00184 (0.00608)
nonsmoker	0.000544 (0.00207)	0.00171 (0.00654)	-0.00114 (0.00434)	-0.00112 (0.00427)
smokerquit	0.00699 (0.00510)	0.0206 (0.0141)	-0.0147 (0.0107)	-0.0129 (0.00852)
alcoholfreq	-0.0124*** (0.00181)	-0.0416*** (0.00636)	0.0257*** (0.00367)	0.0283*** (0.00452)
alcohol_occa	0.00339 (0.00266)	0.0104 (0.00791)	-0.00711 (0.00559)	-0.00664 (0.00499)
exercise	-0.00422 (0.00225)	-0.0137 (0.00758)	0.00879 (0.00466)	0.00915 (0.00517)
nobese	-0.0138** (0.00463)	-0.0388*** (0.0116)	0.0290** (0.00959)	0.0237*** (0.00660)
educ_1	-0.00383 (0.00215)	-0.0123 (0.00703)	0.00800 (0.00448)	0.00812 (0.00471)
educ_2	-0.0159*** (0.00348)	-0.0199* (0.00984)	0.0306*** (0.00928)	0.00517 (0.00634)
educ_3	-0.0143*** (0.00398)	-0.0390*** (0.0115)	0.0490*** (0.0110)	0.00430 (0.00744)
educ_4a	-0.00947* (0.00407)	-0.0327* (0.0154)	0.0195* (0.00810)	0.0227* (0.0114)
single	-0.00890 (0.00602)	0.00902 (0.0154)	-0.0262 (0.0157)	0.0261** (0.00960)
divorce	0.00759 (0.00888)	0.0222 (0.0242)	-0.0159 (0.0186)	-0.0139 (0.0145)
widow	-0.00799 (0.00422)	-0.0377** (0.0139)	0.0555*** (0.0161)	-0.00977 (0.0118)
separate	0.0129 (0.0208)	0.0359 (0.0515)	-0.0269 (0.0430)	-0.0218 (0.0293)
unemp	0.0363*** (0.00697)	0.00882 (0.0133)	-0.0462** (0.0141)	0.00110 (0.00855)
housekeep	0.00834** (0.00263)	0.0247*** (0.00730)	-0.0175** (0.00551)	-0.0155*** (0.00443)
disable	0.268*** (0.0512)	0.0116 (0.0574)	-0.242*** (0.0575)	-0.0379 (0.0339)
retired	0.00406 (0.00320)	0.0124 (0.00944)	-0.00852 (0.00673)	-0.00790 (0.00592)
rehired	-0.0157**	-0.0588*	0.0312**	0.0433*



	(0.00551)	(0.0246)	(0.00963)	(0.0205)
workshift	0.0136*** (0.00406)	0.0380*** (0.0101)	-0.0284*** (0.00841)	-0.0232*** (0.00578)
ag_labor	0.00630** (0.00205)	0.0192** (0.00603)	-0.0132** (0.00429)	-0.0123** (0.00379)
urban_hukou	0.00414 (0.00357)	0.00501 (0.00893)	-0.0209* (0.00891)	0.0117* (0.00581)
male	-0.00169 (0.00404)	-0.00936 (0.0106)	0.0361*** (0.0103)	-0.0250*** (0.00683)
height	-0.000835*** (0.000226)	-0.00334*** (0.000585)	0.000282 (0.000609)	0.00390*** (0.000380)
age	0.00116*** (0.000132)	0.00597*** (0.000342)	-0.00428** (0.000347)	-0.00286*** (0.000226)
hyper	0.0513*** (0.00702)	0.148*** (0.0137)	-0.142*** (0.0139)	-0.0568*** (0.00709)
diabete	0.0991*** (0.0151)	0.157*** (0.0111)	-0.176*** (0.0212)	-0.0799*** (0.00492)
apoplexy	0.208*** (0.0330)	0.0400 (0.0452)	-0.164*** (0.0476)	-0.0840*** (0.0204)
fracture	0.0445*** (0.00605)	0.0995*** (0.00959)	-0.0890*** (0.0110)	-0.0551*** (0.00458)
rltv_income	-0.00954*** (0.00170)	-0.00235 (0.00357)	0.0119*** (0.00359)	-0.0000338 (0.00204)
rltv_BMI	-0.0789*** (0.00652)	-0.248*** (0.0197)	0.165*** (0.0135)	0.162*** (0.0128)
time	0.00767*** (0.00163)	0.0270*** (0.00425)	-0.0354*** (0.00454)	0.000700 (0.00279)
<i>N</i>	22055			

Standard errors in parentheses

\*  $p < 0.05$ , \*\*  $p < 0.01$ , \*\*\*  $p < 0.001$

Table 7. Random effect generalized ordered probit model with objective health measure

	Num_diseases 1 vs. 2-3	Num_diseases 1-2 vs. 3
sleep	0.541*** (0.05)	0.296*** (0.04)
nonsmoker	-0.003 (0.04)	-0.003 (0.04)
smokerquit	-0.188* (0.08)	-0.188* (0.08)
alcoholfreq	-0.022 (0.04)	-0.022 (0.04)
alcohol_occa	-0.126** (0.05)	-0.126** (0.05)
exercise	-0.080 (0.04)	-0.080 (0.04)
nobese	-0.082 (0.09)	0.256*** (0.07)
educ_1	-0.096* (0.04)	-0.096* (0.04)
educ_2	-0.083 (0.05)	-0.083 (0.05)
educ_3	-0.109* (0.05)	-0.109* (0.05)
educ_4a	0.136 (0.13)	-0.083 (0.10)
single	-0.306*** (0.06)	-0.306*** (0.06)
divorce	0.114 (0.16)	0.114 (0.16)
widow	-0.119* (0.06)	-0.119* (0.06)
separate	0.247 (0.37)	0.247 (0.37)
unemp	-0.127* (0.05)	-0.127* (0.05)
housekeep	-0.087 (0.04)	-0.087 (0.04)
disable	-0.840*** (0.16)	-0.840*** (0.16)
retired	-0.289*** (0.06)	-0.445*** (0.05)
rehired	-0.169 (0.13)	-0.169 (0.13)
workshift	-0.158* (0.08)	-0.158* (0.08)

	(0.07)	(0.07)
ag_labor	0.288***	0.097*
	(0.06)	(0.04)
urban_hukou	-0.090*	-0.090*
	(0.04)	(0.04)
male	0.046	0.046
	(0.05)	(0.05)
height	-0.005	-0.010***
	(0.00)	(0.00)
age	-0.018***	-0.030***
	(0.00)	(0.00)
rltv_income	0.003	0.003
	(0.01)	(0.01)
rltv_BMI	-0.902***	-1.300***
	(0.15)	(0.12)
time	0.035	-0.064***
	(0.02)	(0.02)
_cons	4.631***	5.768***
	(0.53)	(0.46)
<hr/>		
rho		
_cons	0.369***	
	(0.02)	
<hr/>		
N	22664	
<hr/>		

Table 8. Marginal effect of random effect Generalized Ordered probit model with objective health measure

	Marginal effects for p(Multi-Disease)	Marginal effects for p(One Disease)	Marginal effects for p(No Disease)
sleep	-0.0317*** (0.00301)	-0.0159** (0.00505)	0.0476*** (0.00566)
nonsmoker	0.000159 (0.00243)	0.000269 (0.00412)	-0.000428 (0.00656)
smokerquit	0.0130* (0.00593)	0.0203* (0.00859)	-0.0333* (0.0145)
alcoholfreq	0.00134 (0.00240)	0.00225 (0.00403)	-0.00359 (0.00643)
alcohol_occa	0.00829* (0.00345)	0.0134* (0.00534)	-0.0217* (0.00877)
exercise	0.00513 (0.00293)	0.00845 (0.00470)	-0.0136 (0.00762)
nobese	0.00473 (0.00502)	-0.0511*** (0.0115)	0.0464*** (0.0131)
educ_1	0.00608* (0.00282)	0.0101* (0.00456)	-0.0161* (0.00737)
educ_2	0.00517 (0.00304)	0.00863 (0.00498)	-0.0138 (0.00801)
educ_3	0.00697 (0.00369)	0.0115 (0.00588)	-0.0184 (0.00957)
educ_4a	-0.00756 (0.00649)	0.0217 (0.0141)	-0.0142 (0.0169)
single	0.0225*** (0.00569)	0.0338*** (0.00758)	-0.0563*** (0.0132)
divorce	-0.00641 (0.00831)	-0.0114 (0.0155)	0.0178 (0.0238)
widow	0.00781 (0.00400)	0.0126* (0.00619)	-0.0205* (0.0102)
separate	-0.0126 (0.0153)	-0.0235 (0.0318)	0.0361 (0.0471)
unemp	0.00834* (0.00375)	0.0135* (0.00579)	-0.0218* (0.00953)
housekeep	0.00554 (0.00298)	0.00912 (0.00477)	-0.0147 (0.00775)
disable	0.0903*** (0.0264)	0.0993*** (0.0194)	-0.190*** (0.0456)
retired	0.0208*** (0.00538)	0.0642*** (0.00906)	-0.0850*** (0.0114)
rehired	0.0116	0.0182	-0.0298

	(0.0102)	(0.0150)	(0.0253)
workshift	0.0107* (0.00490)	0.0170* (0.00732)	-0.0277* (0.0122)
ag_labor	-0.0158*** (0.00311)	0.000319 (0.00556)	0.0155* (0.00614)
urban_hukou	0.00566* (0.00254)	0.00942* (0.00415)	-0.0151* (0.00668)
male	-0.00281 (0.00309)	-0.00477 (0.00524)	0.00759 (0.00833)
height	0.000283 (0.000182)	0.00138*** (0.000301)	-0.00166*** (0.000422)
age	0.00107*** (0.000113)	0.00386*** (0.000190)	-0.00493*** (0.000260)
rltv_income	-0.000205 (0.000826)	-0.000347 (0.00140)	0.000551 (0.00223)
rltv_BMI	0.0549*** (0.00945)	0.158*** (0.0152)	-0.213*** (0.0189)
time	-0.00213 (0.00143)	0.0127*** (0.00238)	-0.0106*** (0.00281)
<i>N</i>	22664		

Standard errors in parentheses

\*  $p < 0.05$ , \*\*  $p < 0.01$ , \*\*\*  $p < 0.001$

Table 9. Regression with SAH and relative income refer to education level subgroups

	Generalized Ordered probit (SAH)		
	1 vs. 2-4	1-2 vs. 3-4	1-3 vs. 4
sleep	0.208*** (0.04)	0.069* (0.03)	0.010 (0.04)
nonsmoker	-0.007 (0.03)	-0.007 (0.03)	-0.007 (0.03)
smokerquit	-0.084 (0.06)	-0.084 (0.06)	-0.084 (0.06)
alcoholfreq	0.164*** (0.02)	0.164*** (0.02)	0.164*** (0.02)
alcohol_occa	-0.040 (0.03)	-0.040 (0.03)	-0.040 (0.03)
exercise	0.052 (0.03)	0.052 (0.03)	0.052 (0.03)
nobese	0.154*** (0.05)	0.154*** (0.05)	0.154*** (0.05)
educ_1	0.053 (0.03)	0.053 (0.03)	0.053 (0.03)
educ_2	0.245*** (0.05)	0.123*** (0.03)	0.021 (0.04)
educ_3	0.269*** (0.06)	0.191*** (0.04)	0.014 (0.04)
educ_4a	0.433*** (0.13)	0.229** (0.07)	0.035 (0.08)
single	0.123 (0.09)	0.001 (0.05)	0.150** (0.05)
divorce	-0.089 (0.10)	-0.089 (0.10)	-0.089 (0.10)
widow	0.111 (0.06)	0.143** (0.05)	-0.055 (0.08)
separate	-0.136 (0.21)	-0.136 (0.21)	-0.136 (0.21)
unemp	-0.354*** (0.06)	-0.130** (0.04)	0.010 (0.05)
housekeep	-0.096*** (0.03)	-0.096*** (0.03)	-0.096*** (0.03)
disable	-1.378*** (0.16)	-0.799*** (0.18)	-0.278 (0.28)
retired	-0.058 (0.04)	-0.058 (0.04)	-0.058 (0.04)
rehired	0.222* (0.10)	0.222* (0.10)	0.222* (0.10)
workshift	-0.156***	-0.156***	-0.156***

	(0.04)	(0.04)	(0.04)
ag_labor	-0.074**	-0.074**	-0.074**
	(0.02)	(0.02)	(0.02)
urban_hukou	-0.057	-0.030	0.081*
	(0.04)	(0.03)	(0.03)
male	0.025	0.036	-0.148***
	(0.05)	(0.04)	(0.04)
height	0.010***	0.012***	0.024***
	(0.00)	(0.00)	(0.00)
age	-0.014***	-0.021***	-0.017***
	(0.00)	(0.00)	(0.00)
hyper	-0.477***	-0.574***	-0.418***
	(0.05)	(0.04)	(0.07)
diabete	-0.739***	-0.739***	-0.739***
	(0.08)	(0.08)	(0.08)
apoplexy	-1.184***	-0.709***	-0.820*
	(0.12)	(0.14)	(0.39)
fracture	-0.415***	-0.415***	-0.415***
	(0.04)	(0.04)	(0.04)
rltv_income2	0.124***	0.050***	0.025
	(0.02)	(0.01)	(0.01)
rltv_BMI	0.968***	0.968***	0.968***
	(0.08)	(0.08)	(0.08)
time	-0.098***	-0.105***	0.006
	(0.02)	(0.01)	(0.02)
_cons	-0.074	-1.491***	-5.449***
	(0.46)	(0.32)	(0.39)
rho			
_cons	0.208***		
	(0.01)		
N	22055		

Standard errors in parentheses

\*  $p < 0.05$ , \*\*  $p < 0.01$ , \*\*\*  $p < 0.001$

Table 10. Regression with objective health measure and relative income refer to education level subgroups

	Num_diseases 1 vs. 2-3	Num_diseases 1-2 vs. 3
sleep	0.542*** (0.05)	0.297*** (0.04)
nonsmoker	-0.002 (0.04)	-0.002 (0.04)
smokerquit	-0.188* (0.08)	-0.188* (0.08)
alcoholfreq	-0.021 (0.04)	-0.021 (0.04)
alcohol_occa	-0.126** (0.05)	-0.126** (0.05)
exercise	-0.078 (0.04)	-0.078 (0.04)
nobese	-0.083 (0.09)	0.256*** (0.07)
educ_1	-0.096* (0.04)	-0.096* (0.04)
educ_2	-0.084 (0.05)	-0.084 (0.05)
educ_3	-0.111* (0.05)	-0.111* (0.05)
educ_4a	0.134 (0.13)	-0.085 (0.10)
single	-0.306*** (0.06)	-0.306*** (0.06)
divorce	0.108 (0.16)	0.108 (0.16)
widow	-0.122* (0.06)	-0.122* (0.06)
separate	0.252 (0.37)	0.252 (0.37)
unemp	-0.132* (0.05)	-0.132* (0.05)
housekeep	-0.090* (0.04)	-0.090* (0.04)
disable	-0.844*** (0.16)	-0.844*** (0.16)
retired	-0.286*** (0.06)	-0.442*** (0.05)
rehired	-0.160 (0.13)	-0.160 (0.13)
workshift	-0.156* (0.04)	-0.156* (0.04)



	(0.07)	(0.07)
ag_labor	0.285***	0.093*
	(0.06)	(0.04)
urban_hukou	-0.088*	-0.088*
	(0.04)	(0.04)
male	0.043	0.043
	(0.05)	(0.05)
height	-0.005	-0.010***
	(0.00)	(0.00)
age	-0.018***	-0.030***
	(0.00)	(0.00)
rltv_income2	-0.016	-0.016
	(0.01)	(0.01)
rltv_BMI	-0.898***	-1.295***
	(0.15)	(0.12)
time	0.035	-0.064***
	(0.02)	(0.02)
_cons	4.627***	5.765***
	(0.53)	(0.46)
<hr/>		
rho		
_cons	0.369***	
	(0.02)	
<hr/>		
<i>N</i>	22664	
<hr/>		

Standard errors in parentheses

\*  $p < 0.05$ , \*\*  $p < 0.01$ , \*\*\*  $p < 0.001$